DEMOCRATIC CONTRACTION? THE POLITICAL CONSEQUENCES OF FELON DISENFRANCHISEMENT IN THE UNITED STATES*

CHRISTOPHER UGGEN, *University of Minnesota* JEFF MANZA, *Northwestern University*

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Universal suffrage is a cornerstone of democratic governance. As levels of criminal punishment have risen in the United States, however, an ever-larger number of citizens have lost the right to vote. We ask whether felon disenfranchisement constitutes a meaningful reversal of the extension of voting rights by considering its political impact. We examine data from legal sources, election studies, and inmate surveys to consider two counterfactual conditions: (1) whether removing disenfranchisement restrictions would have altered the outcomes of U.S. Senate and presidential elections; and, (2) whether applying contemporary rates of disenfranchisement to prior elections would have affected their outcomes. Because felons are drawn disproportionately from the ranks of racial minorities and the poor, disenfranchisement laws tend to take more votes from Democratic than from Republican candidates. We find that felon disenfranchisement played a decisive role in U.S. Senate elections, helping to establish the Republican Senate majority of the 1990s. Moreover, at least one Republican presidential victory would have been reversed even if only former felons had been allowed to vote and at least one Democratic presidential victory would have been jeopardized had contemporary rates of disenfranchisement prevailed during earlier periods.

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The right to vote is a cornerstone of democratic governance and a fundamental element of citizenship in democratic societies, one that "makes all other political rights significant" (Piven and Cloward 2000, p. 2). Although the timing and sequencing of the establishment of formal voting rights varied from country to country, it was almost always a slow, contested and uneven process (e.g.,Rokkan 1970, pp. 31-36; Therborn 1977; Bowles and Gintis 1986, pp. 43-44, 56; Rueschemeyer, Stephens, and Stephens 1992; Collier 1999). As Dahl (1998, p. 89) puts it, "in all democracies and republics throughout twenty-five centuries the rights to engage fully in political life were limited to a minority of adults." Political and economic elites often resisted the extension of voting rights to subordinate groups, including women, youth, the non-propertied, workers, poor people, racial and ethnic groups, and others (Markoff 1996, pp. 45-64; cf. Wiebe [1995] and Keyssar [2000] on the U.S. case).

Yet over the course of the 19th and 20th Centuries, restrictions on the franchise within countries claiming democratic governance have gradually eroded, and universal suffrage has come to be taken for granted as a key component of democracy in both theory and practice (Dahl 1998, p. 90). One recent survey reports that by 1994, fully 96 percent of nation-states claimed to formally enfranchise adult men and women citizens alike (Ramirez, Soysal, and Shanahan 1997, p. 735). To proclaim democratic

¹ To be sure, many of these countries have incomplete or "facade" democracies without fully competitive elections (Markoff 1996, chap. 5). Even within the most democratic

governance today means, at a minimum, universal suffrage for all citizens.

In this paper, we consider a rare and potentially significant counter-example to the universalization of the franchise in democratic societies: restrictions on the voting rights of felons and ex-felons. Felon disenfranchisement constitutes a growing impediment to universal political participation in the United States because of the unusually severe state voting restrictions imposed upon felons and the extraordinarily rapid rise in criminal punishment since the 1970s. While a number of other countries (including the United Kingdom, Russia, and many of the post-Soviet Republics) deny voting rights to prison inmates, the United States is unique in restricting the rights of non-incarcerated felons (who, as we will show below, make up approximately three-quarters of the disenfranchised population). In many European countries, including Ireland, Spain, Sweden, Denmark, and Greece, as well as Australia and South Africa, inmates retain the legal right to vote even while in prison (Fellner and Mauer 1998; Australian Electoral Commission 2001; Ewald 2002). In a number of other countries, voting restrictions are contingent on the length or type of sentence imposed (among these countries are Austria,

countries barriers to participation inevitably persist (e.g., registration requirements, barriers faced by disabled voters, difficulties accessing polling places especially when elections are held on working days). Almost every country excludes non-citizen immigrants from voting (e.g., see Honig 2001).

² We thank Joe Levinson at the Prison Reform Trust, and Femke van der Meulen at the International Centre for Prison Studies, both in London, for making the results of their international survey of felon voting rights in Europe available to us.

Belgium, Italy, and Norway in Europe, and Canada, Australia and New Zealand outside it). Among postindustrial democracies, the United States is virtually the only nation to permanently disenfranchise ex-felons as a class in many jurisdictions, and the only country to limit the rights of individuals convicted of anything other than very rare treason or election-related offenses. Finland and New Zealand disenfranchise some exfelons for specific election offenses, but only for a limited time (Fellner and Mauer 1998). Germany allows judicial discretion to disenfranchise those convicted of election offenses and treason for a maximum of five years beyond their sentence (Demleitner 2000). The U.S. stands alone in the democratic world in imposing restrictions on the voting rights of a very large group of non-incarcerated felons.

As many recent analysts have documented (e.g., Savelsberg 1994; Lynch 1995; Donziger 1996; Sutton 2000), the U.S. is also exceptional for the rate at which it issues felony convictions (and thus removes the right to vote). For example, the 2000 incarceration rate in the U.S. was 686 per 100,000 population, compared to rates of 105 in Canada, 95 in Germany, and only 45 in Japan (Walmsley 2002; U.S. Department of Justice (henceforth USDOJ) 2002; see also Mauer 1997a) and similar disparities can also be found for non-incarcerated felons.

Whether felon disenfranchisement in the United States actually constitutes a threat to democracy, however, is not a simple question. Modern democratic governance entails a set of macro political institutions that register citizens' preferences through (among other things) regular competitive elections (see, e.g., Bollen 1979; Przeworski 1991, chap. 1; Dahl 1998). For democratic governance to be threatened,

disenfranchisement must reach levels sufficient to change election outcomes. Raw counts of the size of the disenfranchised felon population are inconclusive: however much the loss of voting rights matters for affected individuals, there may be no effect on political outcomes and hence, no substantive macro-level impact. Group-level analyses face the same limitations. Some analysts have focused on the disproportionate racial impact of felon disenfranchisement (e.g., Shapiro 1993; Harvey 1994), and the statistical estimate that approximately one in seven African-American men are currently disenfranchised (Fellner and Mauer 1998) has been widely reported. While unquestionably important for many reasons, the disproportionate racial impact of felon disenfranchisement cannot by itself address the implications for American democracy as a whole. In this paper, we develop an appropriate, macro-level test. We suggest that determining whether felon disenfranchisement has had an impact on American democracy requires examining the extent to which it directly alters actual electoral outcomes.

Because felon voting rules are state-specific, the handful of earlier studies of the political consequences of felon disenfranchisement estimated the average impact of disenfranchisement on turnout rates across the states (Miles 2000; Hirschfield 2001). In the analyses developed in this paper, by contrast, we advance an alternative, counterfactual approach that examines specific elections and tests whether the inclusion of felon voters at predicted rates of turnout and party preference would have been sufficient to change the outcomes. We use data on turnout from the Current Population Survey's Voter Supplement Module, and data on voting intention from the National Election Study, to estimate the likely voting behavior of the disenfranchised felon

population. We utilize information on felon characteristics from censuses and surveys of prison inmates to estimate the size and social distribution of the felon population. Combining these data sources, we are able to estimate the net votes lost by Democratic candidates in closely contested U.S. Senate and presidential elections, and to assess the overall impact of felon disenfranchisement on the American political landscape. Finally, we use unique longitudinal data on criminal background and political behavior to test the reasonableness of the assumptions made in our voting analyses, drawing on newly available data from the 2000 wave of the Youth Development Study or YDS (see Mortimer forthcoming).

The paper is in five parts. Part one develops the theoretical and historical background, situating our empirical analyses in the literatures on democratic theory and American criminal justice. The second part describes the logic of our investigation. Part three addresses data sources and methodological issues, presenting our estimates of the size of the disenfranchised felon population in each state. Part four offers two sets of results: estimates of the likely turnout and vote choice of felons, if they had the right to vote, and confirmatory analyses from the YDS. The final part discusses some of the implications of these results.

THEORETICAL AND HISTORICAL BACKGROUND

Models of Universal Suffrage and American Democracy

The current state of democracy in America is frequently characterized as troubled.

Low turnout rates (e.g., Piven and Cloward 2000; Putnam 2000), high levels of public

apathy (Eliasoph 1998), low information and citizen ignorance (Delli Carpini and Keeter 1996), declining trust in the political system (Nye, Zelikow, and King 1997; Brooks and Cheng 2001), a "crisis" of the party system (Burnham 1982) characterized by rising independent partisanship, candidate-centered politics, and voter dealignment (e.g., Wattenberg 1991, 1994) are among the symptoms most frequently identified in the recent literature. Yet, virtually no attention has been paid to issues surrounding the right to vote.³

A lack of attention to voting rights reflects the predominant scholarly consensus that suffrage has been a settled issue since the passage and enforcement of the Voting Rights Act of 1965. Observing the early extension of the franchise to non-propertied white men in the United States in the 1830s, Tocqueville (1969 [1835], p. 59-60) famously asserted that "When a nation begins to modify the elective qualification one can be sure that sooner or later it will abolish it altogether. That is one of the most invariable rules of social behavior." To be sure, democratic governance has been overturned in many countries over the course of the past 150 years, in some cases more than once (e.g., Markoff 1996). ⁴ Such societal-wide democratic reversals have typically entailed the

³ A partial exception to this claim has resulted from the aftermath of the 2000 presidential election and the controversies growing out of the Florida vote. See National Commission on Federal Election Reform (2001).

⁴ Among the most important of these anti-democratic waves were the rise of fascist governments in Europe between the two world wars, and the uneven development of democratic governance in Asia and Central and South America after World War II. See

elimination of democratic institutions and free elections as part of larger shifts to authoritarian forms of governance. In such cases, the right to vote in meaningful elections is either completely eliminated or rendered irrelevant, but *selective* disenfranchisement of particular groups is rarely the source of the turn away from democracy. Democratic theory suggests that suffrage rights are exceptionally sticky: once the vote is extended to a particular segment of the population, it is rarely removed as long as the polity as a whole remains democratic.

The history of suffrage rights in the United States has appeared to many observers to have more or less followed a Tocquevillian model, even if unevenly. Though the struggle to extend the franchise to all continued for some 130 years after Tocqueville wrote, the history of suffrage was generally viewed as a steady march towards universalism (e.g., Williamson 1960; Verba, Nie, and Kim 1978, p. 5; Flanigan and Zingale 2002, pp. 31-34). As keen an observer of the limitations of American democracy as E.E. Schattschneider (1960, p. 100) could assert that "one of the easiest victories of the democratic cause in American history has been the extension of the suffrage...The struggle for the ballot was almost bloodless, almost completely peaceful, and astonishingly easy." The dominant assumption in the literature today is that "at least since the voting rights reforms of the 1960s, political rights have been universalized in the United States. With relatively insignificant exceptions, all adult citizens have the full compliment of political rights" (Verba, Scholzman, and Brady 1995, p. 11).

Recent critical historical accounts have challenged unilinear models of democratic

Rueschemeyer, Stephens, and Stephens (1992) for a global overview.

extension, emphasizing the uneven development of suffrage over the course of American history (e.g., Shklar 1991; Rogers 1992; Wiebe 1995; Keyssar 2000). This newer scholarship describes the halting and at times reversible processes through which universal suffrage finally came to be adopted in the United States. Examinations of state and local variation in the timing and extension of the franchise reveal this pattern most clearly. The possibility that growing felon disenfranchisement may constitute a challenge to the legitimacy of democratic elections, however, has not generally been considered in this literature (see Keyssar [2000, p. 308] for one notable exception).

The widespread consensus around the view that universal suffrage has been attained seems to be driven by a simple but plausible assumption: there is no reason to think that disenfranchisement has any substantive impact on political outcomes, since it affects only a small group of individuals; hence, while it may be an interesting legal or philosophical question, it does not by itself pose an empirical threat to democratic governance. Yet there are reasons to believe that felon disenfranchisement has not had a neutral impact on the American political system.

Racial minorities (Tonry 1995; Kennedy 1997; Mauer 1999) and the poor (Wolfgang, Thornberry, and Figlio 1987; Wilson and Abrahamse 1992; USDOJ 1993; 2000b) are significantly over-represented in the U.S. criminal justice system. We estimate that 1.8 million of the 4.7 million felons and ex-felons currently barred from voting are African Americans (see Appendix Table 2). Because African Americans are overwhelmingly Democratic Party voters (e.g., Huckfeldt and Kohfeld 1989; Tate 1993; Dawson 1994), felon disenfranchisement erodes the Democratic voting base by reducing

the number of eligible African-Americans voters. Moreover, the white felon population is principally composed of poor or working-class offenders (USDOJ 1993; 2000b) who are also likely to be Democratic voters (though not nearly to the same extent as African Americans) (see e.g., Form 1995; Hout, Brooks, and Manza 1995). According to a nationally representative survey of state prison inmates, less than one-third of all state prisoners had completed high school and fewer than half reported annual income of \$10,000 in the year prior to incarceration (USDOJ 1993, p. 3; 2000b). In the Southern states, where disenfranchisement laws tend to be most restrictive, education and income levels are even lower (tables available from authors). For all of these reasons, then, there is at least the possibility that felon disenfranchisement affects the outcomes of democratic elections by taking net votes from the Democratic Party.

Criminal Justice and Felon Disenfranchisement

The possibility that felon disenfranchisement could be influencing recent electoral outcomes is largely tied to changes in the criminal justice regime over the past three decades. For a fifty-year period from the 1920s to the early 1970s, United States incarceration rates fluctuated within a narrow band of approximately 110 prisoners per 100,000 people. The policy consensus accompanying this stability was undergirded by a model of "penological modernism" in which the rehabilitation of offenders was the primary goal of incarceration (e.g., Rothman 1980). Structural elements of the criminal justice system, including probation, parole, and indeterminate sentencing, were designed to reform offenders and reintegrate them into their communities. The model began to

Goldwater (in 1964) and Richard Nixon (in 1968) and other conservative and moderate politicians (such as Nelson Rockefeller in New York) successfully promoted more punitive criminal justice policies (Savelsberg 1994; Jacobs and Helms 1996; Beckett 1997). By the mid-1970s, a rising chorus of conservative scholars, policy analysts, and politicians were advocating punitive strategies of deterrence and incapacitation, dismissing the rehabilitative model as "an anachronism" (Martinson 1974, p. 50; Wilson 1975). These trends continued in the 1980s and 1990s, with the Reagan, Bush, and Clinton Administrations aggressively focusing the nation's attention on problems associated with drug use and the incarceration of drug offenders (see, e.g., Beckett and Sasson 2000).

The success of the conservative crime policy agenda over the past three decades has had a remarkable impact, producing an enormous increase in felony convictions and incarceration, and a corresponding increase in rates of felon disenfranchisement. Since 1970, the number of state and federal prisoners has grown by over 600 percent, from fewer than 200,000 to almost 1.4 million (USDOJ 1973, p. 350; 2001a, p. 1). Other correctional populations have also grown in rate and number, with the number of felony probationers and parolees quadrupling from 1976 to 2000 (USDOJ 1979; 2001b). When jail inmates are added to state and federal prisoners, approximately 2 million Americans are currently incarcerated, with an additional 4.5 million supervised in the community on probation or parole (USDOJ 2000a), and some 9.5 million ex-offenders in the general population (Uggen, Manza and Thompson 2000).

Not all of these felons and ex-felons are disenfranchised. Ballot restrictions for felons are specific to each state. They were first adopted by some states in the post-Revolutionary era, and by the eve of the Civil War some two dozen states had statutes barring felons from voting or felon disenfranchisement provisions in their state constitutions (Keyssar 2000, pp. 62-63; Behrens, Uggen, and Manza 2002). In the post-Reconstruction South, such laws were extended to encompass even minor offenses (Keyssar 2000, p.162), as part of a larger strategy to disenfranchise African Americans that also included devices such as literacy tests, poll taxes, and grandfather clauses (e.g., Kousser 1974). In general, some type of restriction on felon voting rights gradually came to be adopted by almost every state, and at present 48 of the 50 states bar felons – in most cases including those on probation or parole – from voting. At least ten of those states also bar ex-felons from voting, two other states permanently disenfranchise recidivists, and one more state requires a post-release waiting period.⁵

⁵ As of May 2002, Vermont and Maine are the only states that allow incarcerated felons to vote. Referenda eliminated voting rights for Utah and Massachusetts inmates in 1998 and 2000 respectively. Alabama, Florida, Iowa, Kentucky, Mississippi, Nevada, Tennessee (for those convicted prior to 1986), Virginia, Washington (for those convicted prior to 1984), and Wyoming permanently disenfranchise felons unless reinstated by a clemency procedure. Arizona and Maryland permanently disenfranchise certain recidivists (those with two or more felony convictions) and Delaware requires a five-year waiting period. New Mexico rescinded permanent ex-felon disenfranchisement in 2001, and Maryland narrowed its ban on ex-felons in 2002.

Overall, the combination of increasing convictions, laws that prevent most felons from voting, and the steady cumulative growth of the disenfranchised ex-felon population in those states that permanently restrict their voting rights has produced significant overall growth in the disenfranchised population. Our estimates suggest that the total disenfranchised population has risen from less than one percent of the electorate in 1976 to 2.3 percent of the electorate in 2000. Figure 1 shows the steady growth of the proportion of the voting age population disenfranchised over this period. The slight dips in the mid-1970s and late-1990s reflect certain states liberalizing restrictions on ex-felons (see Behrens, Uggen, and Manza 2002; Manza and Uggen forthcoming).

[Figure 2 about here.]

PRIOR RESEARCH AND STRATEGY OF ANALYSIS

Our primary research question in this paper is whether felon disenfranchisement has had meaningful political consequences. Would election outcomes have differed if the disenfranchised had been allowed to vote? To fully answer this counterfactual question, we need to determine how many felons would have turned out to vote, how they would have voted, and whether those choices would have influenced electoral outcomes. If so, a closely related consideration is whether disenfranchisement has affected public policy through feedback processes tied to these electoral outcomes. Figure 2 provides a schematic representation of the questions we pose in this paper. The burden in this paper is to estimate (a), (b), and (c), although in the conclusion we suggest some possible answers regarding (d) as well.

[Figure 2 about here.]

These are difficult questions to answer. A group the size of the disenfranchised felon population *could* have a considerable political impact, but given its composition, neither its rate of political participation nor its preferences are likely to mirror those of the general population. In this case, and in observational research more generally, information is missing on an important counterfactual condition (Holland 1986). If we could assume unit homogeneity, that felons would have voted in the same numbers and with the same preferences as non-felons, we could simply count the disenfranchised felons and apply national turnout and party preference averages. Because felons differ from non-felons in ways that are likely to affect political behavior, however, this sort of blanket assumption is likely untenable.

Another way to measure impact is to estimate the average causal effect of a treatment -- in this case laws stripping criminals of their voting rights. In a state-level analysis of National Election Study data, Miles (2000) reports that rates of voter registration and turnout tend to be lower in states with strict felon disenfranchisement laws than in states lacking such laws, but that the differences are not statistically significant (cf. Hirschfield 2001). Although such studies provide evidence about the statistical significance of the *average* effect of disenfranchisement – and suggest that this average effect is likely to be small – it is possible that even such small differences may have great practical significance.

First, it may be reasonable to examine particular elections rather than overall impact, because political choices are less about average causal effects than about tipping points. In some elections, particularly those in two-party systems requiring a simple

plurality for victory (as in most U.S. elections), a small number of votes is often decisive. In this case, we also have a great deal more information at our disposal than the standard statistical approach assumes, since we have access to population data rather than sample data. We know the precise number of votes cast for each candidate and the plurality or margin of victory in each election. We also know the exact number of prisoners, probationers, and parolees in each state who cannot vote, and we can reasonably estimate the number of ex-felons in states that restrict their voting rights as well. The only real question is how many of them would have turned out to vote, and which candidate they would have selected.

Assuming that nothing else about the candidates or elections would have changed, we therefore undertake a historical accounting of the counterfactual condition: what would have happened had felons been allowed to vote? We calculate the number of felons and ex-felons affected, then estimate turnout and vote choice on the basis of their known characteristics to determine the number of votes lost to Democratic candidates. To assess the political consequences of disenfranchisement, we then compare the actual margin of victory with counterfactual results that take into account the likely political behavior of disenfranchised felons.

DATA AND METHODS

Turnout and Vote Choice

Our analyses of turnout and vote choice utilize standard election data sources. To derive turnout estimates for the disenfranchised population, we analyze data from the

Voter Supplement File of the Current Population Survey (CPS). The CPS is a monthly survey of individuals conducted by the Bureau of the Census. Since 1964, in each November of even-numbered (national election) years, the survey includes questions about political participation. All sampled households are asked, "In any election some people are not able to vote because they are sick or busy or have some other reason, and others do not want to vote. Did [you/another household member] vote in the election on November ?"

Questions of this type produce slightly inflated estimates of turnout in the CPS series, with the inflation factor ranging from a low of 7.5 percent (1968) to a high of 11.1 percent (1988) in presidential elections between 1964 and 1996 (U.S. Bureau of the Census 1998, p. 2). Accordingly, after obtaining estimated turnout percentages for the felon population, we reduce them by a CPS inflation factor, multiplying predicted turnout rates by the ratio of actual to reported turnout for each election. Because turnout is most overreported among better-educated citizens (Silver, Anderson, and Abramson 1986; Bernstein, Chadha, and Montjoy 2001), inflation rates are likely lower among disenfranchised felons than among non-felons, so this procedure is likely to produce

⁶ The use of proxy respondents to report on the voting behavior of others in the household is a potentially greater threat to validity. However, Census Bureau verification tests show that proxy and self-reports were in agreement about 99% of the time in 1984 and 98% of the time in 1992 (U.S. Bureau of the Census 1986, p. 10; 1993). We should also note that the CPS has produced much more reliable turnout estimates than the National Election Study, which typically overestimates turnout by 18 to 25 percent.

conservative estimates.

Our estimates of the expected vote choice of disenfranchised felons are developed using National Election Study (NES) data for 1972 to 2000. We begin in 1972 because it is the first presidential election year for which we have reasonably proximate sociodemographic information about incarcerated felons and because it immediately precedes major increases in incarceration. The NES is the premier source of U.S. voting data, with a rich battery of sociodemographic and attitudinal items and the lengthy timeseries needed for this investigation. The biggest drawback of the NES series is that while it asks respondents how they voted in presidential and congressional elections, there are too few respondents (N < 2500) to permit meaningful state-level analyses.⁷

To analyze the expected turnout and vote choice of disenfranchised felons, we do not have any survey data that asks disenfranchised felons how they would have voted. We can, however, "match" the felon population to the rest of the voting age population to derive such an estimate and then test the reasonableness of this approach with a supplementary survey analysis. Our models of political behavior include sociodemographic attributes that have long been shown in voting research to contribute to turnout and vote choice: gender, race, age, income, labor force status, marital status and education (Wolfinger and Rosenstone 1980; Teixeira 1992; Manza and Brooks 1999,

⁷ It would be possible to obtain state-level data for many elections, such as data collected in recent elections by the Voter News Service. Unfortunately, these surveys generally lack the battery of items needed to match the characteristics of the felon population to those of the survey respondents, and thereby derive estimates of felon voting behavior.

chap. 7). We analyze age and education in years as continuous variables. Income is a continuous variable measured in constant dollars. Labor force status, marital status, gender and race are dichotomies (the latter necessitated by the lack of information about Hispanic voters in the NES series prior to the 1980s). We use similar measures for both the turnout analyses (using CPS data) and vote choice analyses (using NES data). Once we have estimated political participation and party preference equations on the general population, we insert the mean characteristics of disenfranchised felons into these equations to obtain their predicted rates of turnout and Democratic Party preference. We obtain information on the sociodemographic characteristics of convicted felons from the *Survey of State Prison Inmates* data series (e.g., USDOJ 1993; 2000b).

The dependent variables in both the turnout and voting analyses are dichotomous, so we estimate logistic regression models of the probabilities of participation and Democratic vote choice respectively. In the turnout equations, the outcome is coded "1" for voted, and "0" for not voted. In the voting equations, the outcome is coded "1" for Democratic and "0" for Republican choice. We only consider major party voters, as in Senate elections very few third party or independent candidates have come close to winning office. 9 Coefficients from these regressions are reported in Appendix Table 3.

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⁸ Ideally, we would also have data on partisanship, and candidate, and policy preferences to develop estimates of felons' voting behavior. Because such information is currently unavailable, we supplement the national analysis with a community study below.

⁹ The only independent candidate to win a Senate seat since 1972 was Harry F. Byrd, Jr. of Virginia in 1976.

Legal Status and Correctional Populations

In addition to estimating the likelihood of voting and the partisan alignment of felons, we must also determine their absolute numbers in each state. To establish which correctional populations to count among the disenfranchised population, we examined the elector qualifications and consequences of a felony conviction as specified in state constitutions and statutes (see Manza and Uggen forthcoming) and referenced secondary sources detailing the voting rights of offenders (e.g., Burton, Cullen, and Travis 1986; Olivares, Burton, and Cullen 1996; USDOJ 1996; Mauer 1997b; Fellner and Mauer 1998; Allard and Mauer 1999). To establish the number of disenfranchised felons currently under supervision, we sum the relevant prison, parole, felony-probation, and convicted felony jail populations. The felons under supervision data come from Justice Department publications, such as the Correctional Populations in the United States series. We estimate that 3 million current felons were legally disenfranchised on December 31, 2000 or slightly less than half of the 6.5 million adults under correctional supervision (USDOJ 2001b). For most states, this is a rather straightforward accounting of the prison, parole and felony probation populations. 10 Convicted felons who serve their sentences in jail

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¹⁰ Connecticut, Rhode Island, Vermont, Delaware, Alaska, and Hawaii combine their prison and jail systems. In such cases, we classify felons serving greater than one year as prisoners and felons with shorter sentences as jail inmates (taking 10 percent of the latter group to represent convicted felony jail inmates). For five states that do not distinguish felony and non-felony probationers, we estimate that 50 percent of probationers are

represent a smaller but potentially important group not considered in prior estimates (Mauer 1997b). In 1998, for example, 24 percent of felony convictions resulted in jail sentences (USDOJ 1998). We therefore include a conservative estimate of the number of convicted felons in jail – 10 percent of the total jail population.

These "head counts" are based on excellent data by social scientific standards. Estimating the number of disenfranchised *ex*-felons not currently under supervision, however, is a greater challenge. Existing estimates vary widely with the assumptions made by researchers. Important early work by the Sentencing Project (Mauer 1997b; Fellner and Mauer 1998) based estimates on national felony conviction data and state-level reports of criminal offenses between 1970 and 1995. While valuable, such procedures may make untenable stability and homogeneity assumptions, such as applying national information on racial composition and criminal convictions to individual states. Moreover, they do not account for the mortality of deceased felons or consider those convicted prior to 1970 or after 1995.

We develop alternative estimates based on exits from (rather than entry into) correctional supervision. Our data sources include the annual Sourcebook of Criminal Justice Statistics and Correctional Populations series, Probation and Parole in the United States, and Prison and Jail Inmates at Midyear. For early years, we also referenced National Prisoner Statistics, and Race of Prisoners Admitted to State and

felons (a more conservative figure than the 52 percent national average) (USDOJ 2001b). Jail figures for 2000 were estimated by applying state-specific prison growth rates from 1999 to 2000 to 2000 jail populations.

Federal Institutions, 1926-1986 (all USDOJ publications). We determine the median age of released prisoners based on annual data from the National Corrections Reporting Program (USDOJ 1983-1996). We use recidivism data from national probability samples of prison releasees (USDOJ 1989) and probationers (USDOJ 1992) to establish the number reincarcerated. We then compile life tables for the period 1948-2000 to determine the number of released felons lost to recidivism (and therefore already included in our annual head counts) and mortality each year (see, e.g., Bonczar and Beck 1997). Each cohort of disenfranchised releasees is thus successively reduced each year and joined by a new cohort of releasees. This allows us to compute the number of exfelons no longer under correctional supervision for states that disenfranchise ex-felons.

Our recidivism estimates are based on Justice Department studies of prisoners (1986) and probationers (1992). The prisoner and parolee reincarceration rate is 18.6 percent at one year, 32.8 percent at two years, and 41.4 percent at three years. For probationers and jail inmates, the corresponding three-year failure rate is 36 percent. To extend the analysis to subsequent years, we computed a trend line based on the ratio of increases in Hoffman and Stone-Meierhoefer's (1980) study of federal prisoners. By year ten, we estimate a 59.4 percent recidivism rate among former prisoners and parolees, which increases to 65.7 percent by year 52 (the maximum duration in the analysis). Because these rates exceed those of most long-term recidivism studies, they should yield conservative estimates of the disenfranchised ex-felon voting base. We calculate mortality based on the expected number of deaths for African American males (the group with the highest mortality rates) at the median age of release for each state, multiplied by

a constant factor of 1.46 to match the high death rates observed in the recidivism study (USDOJ 1986). Age- and year-specific mortality data were obtained from the *Statistical Abstract* series ("Expectation of Life and Expected Deaths, by Race, Sex, and Age").

These ex-felon estimates also account for the fact that some states restore the civil rights of many releasees or only disenfranchise certain ex-felons. Florida, for example, has restored voting rights to over 160,000 disenfranchised felons since the 1960s and does not impose felony adjudication for some probationers who successfully complete their sentences.

THE POLITICAL IMPACT OF FELON DISENFRANCHISEMENT

Turnout and Party Preference

Table 1 shows the estimated national participation rates and voting preferences for disenfranchised felons by year since 1972. These estimates are based on the voting behavior of those matching felons in terms of gender, race, age, income, labor force status, marital status, and education, reduced for overreporting in the CPS. In short, they provide evidence regarding the likely behavior of hypothetical felon and ex-felon voters. Our estimates of felon turnout range from a low of 20.5 percent (for the 1974 Congressional elections) to a high of 39 percent (for the 1992 presidential election). On average, we predict that about 35 percent of disenfranchised felons would have turned out to vote in presidential elections, and that about 24 percent would have participated in Senate elections during non-presidential election years. Although these numbers are well below the corresponding rates among non-felons, they suggest that a non-trivial

proportion of disenfranchised felons were likely to have voted if permitted to do so.

[Table 1 about here.]

According to our analysis of party choice in Table 1, our hypothetical felon voters showed strong Democratic preferences in both presidential and senatorial elections. In recent presidential elections, even comparatively unpopular Democratic candidates such as George McGovern in 1972 would have garnered almost 70 percent of the felon vote. These Democratic preferences are less pronounced and somewhat less stable in senatorial elections. Nevertheless the survey data suggest that Democratic candidates would have received about seven of every ten votes cast by the felons and ex-felons in 14 of the last 15 U.S. Senate election years. By removing those with Democratic preferences from the pool of eligible voters, felon disenfranchisement has provided a small but clear advantage to Republican candidates in every presidential and senatorial election from 1972 to 2000.

Impact on Individual U.S. Senate Elections

We next use these turnout and party preference rates to gauge the impact of felon disenfranchisement on U.S. presidential and Senate elections. We obtained information on victory margins and Senate composition from standard election data sources (Congressional Quarterly's *America Votes* bienniel series). Table 2 applies the voting behavior estimates from Table 1 to these election data, identifying seven elections that may have been overturned if disenfranchised felons had been allowed to participate.

[Table 2 about here]

To determine the *net* Democratic votes lost to disenfranchisement, we first

multiply the number of disenfranchised felons by their estimated turnout rate (in each state), and the probability of selecting the Democratic candidate. 11 Since some felons would have chosen Republican candidates, we then deduct from this figure the number of Republican votes lost to disenfranchisement, which we obtain in a similar manner. For the 1978 Virginia election detailed in the top row of Table 2, for example, we estimate that 15,343 of the state's 93,554 disenfranchised felons would have voted (16.4 percent). We further estimate that 12,305 of these voters would have selected Andrew Miller, the Democratic candidate (80.2 percent of 15,343), and that the remaining 19.8 percent (or 3,038) would have chosen John Warner, the Republican candidate. This results in a net total of 9,268 Democratic votes lost to disenfranchisement in the 1978 U.S. Senate race in Virginia, almost double the actual Republican victory margin of 4,721 votes.

In recent policy debates over felon disenfranchisement, restoring voting rights has been most widely discussed for ex-felons who have completed their sentences (see, e.g., Sengupta 2000; Bush 2001). Yet some analysts have asserted that ex-felon voting restrictions are "electorally insignificant" (Harvard Law Review Note 1989, p. 1303). Is this assumption accurate? The results in Table 2 offer a new perspective on this issue. Recall that most states only deprive those currently under some form of correctional supervision of the right to vote, with only 15 states additionally disfranchising some or all ex-felons in 2000 (see Appendix Table 1). In only one instance (the late Paul Coverdell's election in Georgia in 1992), however, was a Senate election likely to have been

¹¹ We draw on the large CPS sample to derive state-level turnout estimates for these key races.

overturned as a result of the disenfranchisement of those actively under correctional supervision. 12 Even in this case, however, the number of current prisoners (25,290) and convicted felony jail inmates (2,163) was too small to affect the election. Rather, it was the large number of felony probationers (80,639, or a full 61 percent of the state's disenfranchised population) and parolees (23,819, or 18 percent of disenfranchised Georgians) that likely cost the Democrats the election. As this case illustrates, the political impact varies with the particular correctional populations disenfranchised. The other reversible cases all include net Democratic vote losses from *ex-felon* voters.

Impact on U.S. Senate Composition

Would changes to a handful of elections have had any real impact? Since 1978, there have been over 400 Senate elections, and we find 7 outcomes that may have been reversed if not for the disenfranchisement of felons and ex-felons. Yet even this small number might have shifted the balance of power in the Senate, which has been fairly evenly divided between the two major parties over this period. To assess this possibility, we recomputed the U.S. Senate composition after each election. Since two Republican seats were overturned in the 1978 elections, the Democratic majority would have increased from 58:41 to 60:39. We followed the beneficiaries of these closely contested elections to see how long their seats remained under Republican control. John Warner of Virginia remains in office today and John Tower's Texas seat also remains in Republican

¹² Georgia's state constitution disenfranchised "until the granting of pardon" until 1983, when the constitutional ban was lifted upon "completion of this sentence."

hands (with Phil Gramm holding office in 2002). Although we cannot know whether the Democratic Party would have held these seats in subsequent elections, the well-known advantages of incumbency make this a plausible scenario. Of the 32 U.S. Senate elections in 1978, the incumbent party retained the seat through at least 1990 in 29 cases (91 percent), through at least 1996 in 27 cases (84 percent) and through at least 2002 in 23 cases (72 percent). Because incumbent parties are unlikely to hold such seats indefinitely, we cumulate the counterfactual using the more reasonable (though untested) assumption that the Democrats would have retained these seats as long as the Republicans who narrowly defeated them. This procedure makes strong *ceteris paribus* assumptions, however, so Table 2 also shows "limited counterfactual" results that assume the victor's party would lose the seat immediately after a single six-year term.

After the 1984 elections, the Republicans held a narrow 53:47 Senate majority. Under the cumulated counterfactual scenario in which disenfranchised felons had voted, the Democrats may have achieved parity with the Republicans. In the Kentucky election of 1984, the Republican candidate (Mitch McConnell) narrowly defeated the Democratic nominee by 5,269 votes. Since Kentucky disenfranchises ex-felons as well as current inmates, parolees, and felony probationers, the total number disenfranchised was over 75,000 in 1984. Because 1984 was a presidential election year, turnout was relatively high, and our voting preference model indicates that almost 70% of the felon voters would have selected the Democratic candidate. Thus, almost 11,000 Democratic votes were likely lost to disenfranchisement in this election, more than twice the 5,269-vote Republican plurality. With the addition of this seat and the Virginia and Texas seats

discussed above, the counterfactual Senate composition shows an even 50:50 party distribution.

Pursuing the counterfactual to the present day, we find that Democratic candidates are likely to have prevailed in Florida (1988), Georgia (1992), and in Kentucky's other seat (1998) if felons had been allowed to vote, with a narrower reversal occurring in Wyoming (1988). Without felon disenfranchisement, our cumulative counterfactual suggests that Democrats may well have controlled the Senate throughout the 1990s. Although it is possible that both parties may have shifted course or that other factors would have arisen to neutralize this impact, it seems likely that the Senate deadlock after the 2000 elections would have been broken in favor of the Democrats if the ballot had been returned to disenfranchised felons. We discuss the implications of these shifts in the conclusion.

Further Tests

Our counterfactual results are startling, but subject to a number of assumptions that might be challenged. How robust are they? On the one hand, our estimates of disenfranchised felon turnout are based on sociodemographic characteristics at the time of incarceration. For the ex-felons who represent more than one third of the entire disenfranchised population, we are likely to *understate* political participation because our sociodemographic measures are taken at the time of incarceration. That is, they do not consider changes in age and personal circumstances (for example, greater residential stability, labor force attachment, and marriage) linked to turnout. During or after

completion of their sentences, many (though certainly not all) ex-felons will acquire greater education and more stable attachments to work, family, and their communities (Sampson and Laub 1993) that may conceivably increase their likelihood of voting.

Moreover, the surveyed *inmate* population is generally less educated, less likely to be married, and less likely to be employed than the entire *felon* population, which also includes a large number of felony probationers who were never sent to prison. For these reasons, we might expect felons and ex-felons to be closer to the national turnout mean than suggested by our model, which is based on sociodemographic characteristics at the time of incarceration. If this were the case, higher estimated turnout rates would increase the impact on electoral outcomes. Finally, our estimates only count ten percent of the total jail population among the disenfranchised. Although jail inmates serving time for misdemeanor offenses and those being held prior to trial are legally eligible to vote, they lack access to a polling place, rendering them practically – if not legally – disenfranchised. If we had included all 621,149 jail inmates in 2000 (USDOJ 2001a) among the disenfranchised, the political impact would have been even greater.¹³

But other unmeasured characteristics of felons and ex-felons, beyond those captured by the individual- and group-level sociodemographic information available in inmate surveys, could serve to significantly depress political participation among this group. Felons may be less cognizant of, or less willing to accept, basic norms of

¹³ Absentee ballots are not routinely available in jails, though there have been scattered efforts to register jail inmates in recent elections (e.g., American Civil Liberties Union 2000; Mitchell 2002).

citizenship and acceptable behavior than non-felons with otherwise identical characteristics (Gottfredson and Hirschi 1990). If so, they may be less likely to vote than a model based solely on sociodemographic traits would predict.

Our counterfactual analysis hinges on the key assumption that the political behavior of disenfranchised felons would approximate that of non-felons matched to them in terms of age, race, gender, education, income, and marital status. Although we cannot provide a conclusive test of this assumption, we gathered new data to examine how experiences with the criminal justice system affect voting behavior. The Youth Development Study (YDS) is a longitudinal survey begun among a sample of 9th graders in 1988 in St. Paul (Minnesota) Public Schools (Mortimer forthcoming). By 1998, when most respondents were 24-25 years old, approximately 23 percent had been arrested and 7 percent had been incarcerated. We estimated logistic regression models to see whether a bivariate association exists between criminal justice experiences and voting and, if so, how much of the observed association is due to the socioeconomic and demographic characteristics that we account for in the models presented above.

[Table 3 about here]

Table 3 shows the effects of arrest on turnout and party preference (results for the jail analysis are similar, though there are far fewer jail inmates than arrestees; tables available from authors). As expected, Model 1 shows a significant bivariate relationship between arrest and turnout in the 1996 presidential election: the odds of voting are only about half as high for arrestees as for non-arrestees ($e^{-.681} = .51$). Model 2, however, shows that this effect is reduced to non-significance once race, gender, education,

income, employment, and marital status are included in the full voting behavior model. When these independent variables are set to their mean values, the predicted probability of voting in Model 2 is about 63 percent for arrestees and 69 percent for non-arrestees. It is likely that at least part of this remaining turnout gap is attributable to the legal disenfranchisement of arrestees still under correctional supervision. In Minnesota, those convicted of felonies may not vote until they are "off paper," i.e. have completed probation or parole supervision in addition to any prison sentence. Unfortunately, we cannot determine from these data whether individual arrestees were legally eligible to vote at the time of the election. Model 3 disaggregates the arrest data, showing that those who had been arrested for violent offenses were significantly less likely to vote in 1996. Those convicted of violent offenses are most likely to face long sentences, so a portion of this effect may again be due to legal disenfranchisement. Though the coefficient for violent arrests remains large in magnitude in Model 4, it is again reduced to non-significance when the full set of voting predictors is introduced.

The remaining models in Table 3 predict party preferences in the 1996 presidential and 1998 Minnesota gubernatorial elections. Unlike the turnout models, there is some evidence that criminal justice sanctions remain associated with party preferences even net of our set of voting predictors. In particular, those arrested for drug or alcohol-related offenses were significantly more likely to favor the Democratic presidential candidate Bill Clinton in 1996 and the Independent Party gubernatorial candidate Jesse Ventura in 1998. Although YDS arrestees and jail inmates may not be representative of the U.S. felon population, results from this Minnesota cohort of young adults help to

establish the plausibility of our turnout and party preference models and our inferences regarding the political impact of felon disenfranchisement.

Impact on Presidential Elections

Although the outcome of the extraordinarily close 2000 presidential election could have been altered by a large number of factors, it would almost certainly have been reversed had voting rights been extended to any category of disenfranchised felons.

While Al Gore won a plurality of the popular vote, defeating the Republican George W. Bush by over 500,000 votes, he lost narrowly in the Electoral College. Had disenfranchised felons been permitted to vote, we estimate that Gore's margin of victory in the popular vote would have surpassed one million votes, as shown in Table 4. Regardless of the popular vote, however, one state – Florida – held the balance of power. If disenfranchised felons in Florida had been permitted to vote, the Democrat Gore would certainly have carried the state and the election.

[Table 4 about here.]

As Appendix Table 1 shows, there are more disenfranchised felons in Florida, approximately 827,000, than in any other state. Had they participated in the election at our estimated rate of Florida turnout (27.2 percent) and Democratic preference (68.9 percent), Gore would have carried the state by more than 80,000 votes. As a test on the sensitivity of these results, we halved the estimated turnout rate and consider only *ex*-felons in Table 4. Under the reduced turnout scenario, the Democratic Party's margin of victory is over 40,000 votes. More interesting, perhaps, is the finding reported in Table 4

that even if only ex-felons had been enfranchised in Florida, they would have yielded an additional 60,000 net votes for Gore, more than enough to overwhelm Bush's narrow victory margin (and to reverse the outcome in the Electoral College). And even if we halve the estimated turnout rate, Gore's margin of victory would have exceeded 30,000 votes. We can thus conclude that the outcome of the 2000 presidential race hinged on the narrower question of *ex*-felon disenfranchisement rather than the broader question of voting restrictions on felons currently under supervision.

What about earlier presidential elections? Here we examine a much different counterfactual condition. Since a greater share of the voting age population is disenfranchised now than ever before, some closely contested Democratic political victories of the recent past might have gone to the Republicans had contemporary rates of disenfranchisement prevailed at the time. In particular, two Democratic presidential victories in the last forty years (1960 and 1976) were decided by very narrow margins that might have been threatened under current levels of incarceration and disenfranchisement.

John F. Kennedy won the 1960 presidential election by a popular vote margin of 118,550 and a 303:219 margin in the Electoral College. Had contemporary rates of criminal punishment held at the time, however, it is very likely that Richard M. Nixon would have won the popular vote. As Appendix Table 1 shows, about 4.7 million citizens or 2.28 percent of the voting age population were disenfranchised in 2000 due to felony convictions. If this proportion had held in 1960, about 2.5 million voters would have been disenfranchised, as shown in Table 4 (2.28154 percent multiplied by the voting age

population of 109,672,000). Since the population proportion of convicted felons was actually much lower in 1960 than today, however, we estimate that only about 1.4 million were actually disenfranchised at the time of the 1960 election. ¹⁴ Therefore, at current rates of disenfranchisement, over one million additional citizens would have been denied the vote in 1960. If 40 percent of these new felons had voted (in an election where the overall turnout rate reached a post-World War II peak of 62.8 percent), and 75 percent of this group selected the Democratic candidate, figures in line with our findings for other presidential elections, then Kennedy would have lost approximately 225,000 votes – almost twice the popular vote margin in the election. If the felon turnout rate had been only 20 percent, we find that at current disenfranchisement levels Kennedy would have prevailed by only 6,000 votes. In applying the counterfactual to the Electoral College, our analysis suggests that Nixon would likely have been victorious in New Mexico (with 4 electoral votes) but would have lost by very narrow margins in other states. Therefore, if current rates of disenfranchisement had held in 1960, it is likely that Nixon may have beaten Kennedy in the popular vote, but unlikely that he would have surpassed his electoral vote total.

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¹⁴ Many states altered their disfranchisement regimes between 1960 and 2000 (Behrens, Uggen, and Manza 2002; Uggen and Manza forthcoming) and the 1960 figures account for these legal changes within the limitations of the available data. Prison, parole, and jail information are available for 1960, but probation figures are imputed based on state-specific ratios of probation to other correctional populations. Ex-felon figures are based on releases from 1948 to 1960 only, so they may be understated relative to recent years.

It is doubtful that applying contemporary disenfranchisement rates would have overturned the 1976 election, although the Democratic victory margin would have been considerably narrower. At current rates of disenfranchisement, about 2.5 million additional citizens would have been denied the vote in 1976. Our NES estimates suggest that 34.3 percent of these would have voted and that 80.7 percent of this group would have selected the Democratic candidate Jimmy Carter. This would have accounted for about 525,000 votes, or about 31 percent of Carter's final 1,682,970-vote victory margin. 15

DISCUSSION AND IMPLICATIONS FOR AMERICAN DEMOCRACY

We find that felon disenfranchisement laws, combined with high rates of criminal punishment, may have altered the outcome of seven recent U.S. Senate elections and at least one presidential election. One startling implication of these findings relates to

The NES does not ask respondents how they voted in specific gubernatorial or other state elections, so we cannot model voting behavior in such elections. We can, however, make some informed assumptions to estimate the effect of felon disenfranchisement in gubernatorial elections. If we apply the mean rate of turnout (24 percent) and Democratic preference (73 percent) in Senate elections to these races, it is likely that at least three Republican gubernatorial victories would have been overturned: in Alabama (with James Folsom (D) defeating James Forrest (R) in 1994), New Jersey (James Florio (D) defeating Thomas Kean (R) in 1981), and Texas (John Hill (D) defeating William Clements (R) in 1978).

control over the Senate. Assuming that Democrats who might have been elected in the absence of felon disenfranchisement had held their seats as long as the Republicans who narrowly defeated them, the Democratic Party would have gained parity in 1984 and held majority control of the U.S. Senate from 1986 to the present. Changing partisan control of the Senate would have had a number of important policy consequences: in particular, it might have enabled the Clinton Administration to gain approval for a much higher proportion of its federal judicial nominees, and key Senate committees would have shifted from Republican to Democratic control.

In examining presidential elections, we find that the Republican presidential victory of 2000 would have been reversed had felons been allowed to vote, and that the Democratic presidential victory of 1960 may have been jeopardized had contemporary rates of disenfranchisement prevailed at that time. Disenfranchised felons and ex-felons currently make up 2.28 percent of the voting age population, a figure that we project may rise to 3 percent within 10 years (Manza and Uggen forthcoming; estimates also available upon request). Because the margin of victory in 3 of the last 10 presidential elections has been 1.1 percent of the voting age population or less, felon disenfranchisement could be a decisive factor in future presidential races.

One potentially important implication of these results concerns the differing correctional populations affected by ballot restrictions. We estimate that the disenfranchised population is composed of approximately 35 percent ex-felons, 28 percent probationers, 9 percent parolees, but only 27 percent prison and jail inmates (Manza and Uggen forthcoming). Disenfranchisement of prisoners alone is therefore

unlikely to alter elections, but the numbers mount when those supervised in the community are added, and reach a critical mass in states that disenfranchise ex-felons. Thus, the impact of felon disenfranchisement would have been greatly reduced had exfelons, probationers, and parolees been permitted to vote in all states. Moreover, the philosophical rationale for disenfranchisement, founded on the liberal legal model and Enlightenment conceptions of the social contract, would appear to be much stronger for current prison inmates than for those who have completed their sentence (ex-felons), or those otherwise deemed fit to maintain community ties (probationers and parolees). Just as disenfranchisement is a powerful symbol of felons' diminished civil rights, restoration of voting rights provides a clear marker of reintegration and acceptance as a stakeholder in a community of law-abiding citizens. Although the public opinion evidence is limited, our recent experimental national survey (Manza, Brooks and Uggen 2002) suggests that significant majorities of survey respondents believe that an offender's right to vote should be restored upon release from prison.

Although these results are striking, do they signal a true democratic reversal in the United States? Figure 3 presents data placing felon disenfranchisement in historical context, showing the percentage of states holding felon disenfranchisement provisions from the late 18th century to present. Most states began to restrict the ballot for felons in the mid-19th century, and there is evidence in some states that lawmakers fully appreciated the partisan consequences of their actions (McMillan 1955; Keyssar 2000; Behrens, Uggen and Manza 2002; Manza and Uggen forthcoming). Few states rescinded such measures following the enfranchisement of African-American males (with passage

of the 14th and 15th amendments to the U.S. Constitution) and women (with passage of the 19th amendment). Nor was felon disenfranchisement dismantled during passage of the Civil Rights Act of 1964, Voting Rights Act of 1965, or Voter Registration Act of 1993. Although several states have removed voting restrictions on *ex*-felons since the 1960s (including New Mexico in 2001), most continue to disenfranchise prisoners, probationers, and parolees today. In fact, as Figure 3 shows, a greater percentage of states disenfranchised felons in 2000 than in any prior year.

[Figure 3 about here.]

Today, high rates of criminal punishment, rather than new laws, account for the political impact of felon disenfranchisement. In light of past theory and research on the extension and universalization of suffrage, however, the persistence and expansion of these ballot restrictions are noteworthy. We have shown that about 4.7 million adult U.S. citizens do not enjoy the full compliment of political rights. As the number of disenfranchised felons expands, the electorate contracts. Because the contracted electorate now produces different political outcomes than the full citizenry, mass incarceration and felon disenfranchisement have clearly impeded and perhaps reversed the historic extension of voting rights. Nevertheless, we must also note a number of caveats to these findings. First, our counterfactual examples rely upon a *ceteris paribus* assumption – that nothing else about the candidates or elections would change save the voting rights of felons and ex-felons. Had these laws changed, other forces might have arisen to negate the political influence of felons and ex-felons. Moreover, although the Democrats lose votes to felon disenfranchisement, they may also have gained votes by

attempting to be just as punitive as Republicans. 16

Second, our estimated vote choice and turnout analysis matched non-felons to felons on the basis of region, gender, race, age, labor force status, marital status and education. Although non-felon voters resemble felons in many respects, we cannot be certain that the experience of criminal conviction itself may not suppress, (or conversely, mobilize) political participation. Our analysis of new survey data on this question provides some reassurance that our turnout and party preference estimates are reasonable, though the YDS results do not constitute a conclusive test of the effects of felony convictions on political behavior. Third, our analyses have assumed that felon disenfranchisement laws are well enforced, and that felons and ex-felons do not attempt to vote in disregard of these laws. Surely some disenfranchised felons have cast ballots, although occasional charges of fraudulent voting have not, upon further investigation, produced significant evidence of illegal voting. There is also some evidence that state

¹⁶ By embracing a law and order agenda in the 1990s, Democrats have neutralized crime as a partisan political issue (e.g., Poveda 1994). Research decomposing the unique contribution of crime policy to individual vote choice is needed to determine whether the votes gained by such strategies outweigh the votes lost with the disenfranchisement of potential Democratic voters. We should note, however, that returning the ballot to felons is not necessarily inconsistent with a crime control agenda. One may advocate extending the franchise on public safety and reintegrative grounds, arguing that ex-felons who become stakeholders in their communities will have lower rates of recidivism.

authorities have improperly purged ex-felons from the rolls, thereby offsetting or perhaps eclipsing the number of votes cast fraudulently (cf. Palast 2000).

Despite these important caveats, we find considerable evidence that ballot restrictions for felons and ex-felons have had a demonstrable impact on national elections, and in this sense rising levels of felon disenfranchisement constitute a reversal of the universalization of the right to vote. Further, our focus on national and state-level elections understates the full impact of felon disenfranchisement. Because of the geographic concentration of disenfranchised felons and ex-felons in urban areas, it is likely that such impact is even more pronounced in local or district-level elections, such as House, state legislative, and mayoral races. ¹⁷ Moreover, our analysis has only examined past elections. Unless disenfranchisement laws change, the political impact is likely to intensify in the future. Even if the numbers of those incarcerated begin to level off (USDOJ 2001a), the number of disenfranchised ex-felons will continue to rise for several years in those states that restrict their franchise.

Although we have specified the political consequences of felon disenfranchisement, we have only touched on the origins of these laws and the mass incarceration phenomenon that gives such force to them today. These questions are important for situating felon disenfranchisement within a broader model of social control

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¹⁷ We should note that in many local races, especially in mostly black urban districts, the partisan impact of felon disenfranchisement might be diminished because Republican candidates are already uncompetitive in these districts.

of dispossessed groups. Proponents of the "new penology" argue that the focus of criminological interest has recently shifted from the rehabilitation of individual offenders to the social control of aggregate groups (Feeley and Simon 1992; Wacquant 2001). The correctional population is subject to a number of exclusions: they are often ineligible for federal grants for education (such as Pell Grants), they have restricted access to social programs, they face sharp disadvantages in the labor market (Western and Beckett 1999), and they must live with the social stigma associated with a felony conviction. Restricted access to the ballot is but a piece of a larger pattern of social exclusion for America's vast correctional population.

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Table 1. Estimated Turnout and Voting Preferences of Disenfranchised Felons

	Pr	esidential E	Elections	Sena	ate Elections
Year	Candidate	Turnout	% Democratic	Turnout	% Democratic
1972	McGovern	38.2%	69.1%	38.2%	68.2
1974				20.5	77.1
1976	Carter	34.3	80.7	34.3	79.6
1978				23.0	80.2
1980	Carter	35.7	66.5	35.7	69.6
1982				26.2	76.8
1984	Mondale	38.2	70.1	38.2	68.9
1986				25.3	73.6
1988	Dukakis	30.0	72.8	30.0	79.4
1990				23.8	80.5
1992	Clinton	39.0	73.6	39.0	74.7
1994				23.1	52.2
1996	Clinton	36.1	85.4	36.1	80.4
1998				23.9	69.7
2000	Gore	29.7	68.9	29.7	76.1

Table 2. The Impact of Felon Disenfranchisement on U.S. Senate Elections 1978-2000

'		Disenfran	chised Po	pulation	Estimated Voting Behavior			Republic	can Victor	ry Margin	Senate Composition			
							Net Dem.		Counter	Rep. Held		Limited	Cumulated	
		Current	Ex-		Turnout	Percent	Votes	Actual	-factual	Seat	1	Counter-	Counter-	
Year	State	Felons	Felons	Total	Rate	Dem.	Lost	Margin	Margin	Through	Actual ¹	factual	factual	
1978	Virginia ²	21,776	71,788	93,564	16.4%	80.2%	9,268	4,721	-4,547	2002+	58:41-D	60:39-D	60:39-D	
1978	Texas ³	100,707	89,662	190,369	13.4%	80.2%	15,408	12,227	-3,181	2002+	58:41 - D	60:39-D	60:39-D	
1980	unchanged	!									53:46-R	51:48-R	51:48-R	
1982	unchanged	!									54:46-R	52:48-R	52:48-R	
1984	Kentucky ⁴	20,583	54,481	75,064	38.5%	68.9%	10,925	5,269	-5,655	2002+	53:47-R	52:48-R	50:50	
1986	unchanged	!									55:45 - D	56:44-D	58:42-D	
1988	Florida ⁵	87,264	206,247	293,512	26.5%	79.4%	45,735	34,518	-11,217	2000	55:45-D	58:42-D	60:40-D	
1988	Wyoming ⁶	3,013	6,969	9,982	24.5%	79.4%	1,438	1,322	-116	2006+	55:45 - D	58:42-D	60:40-D	
1990	unchanged	!									56:44-D	58:42-D	61:39-D	
1992	Georgia ⁷	131,911	0	131,911	29.6%	74.7%	19,289	16,237	-3,052	2000	57:43-D	60:40 - D	63:37-D	
1994	unchanged	!									52:48-R	51:49-R	54:46-D	
1996	unchanged	!									55:45-R	54:46-R	51:49-D	
1998	Kentucky ⁸	31,456	94,584	126,040	25.4%	69.7%	12,614	6,766	-5,848	2004+	55:45-R	54:46-R	50:50-D	
2000	unchanged	!									50:50	51:49-D	55:45-D	
17.	•	•			•		•				•	•	-	

Notes:

¹ Data on actual senate composition taken from *Senate Statistics: Majority and Minority Parties*.

² In Virginia, Warner (R) defeated Miller (D) in 1978, Harrison in 1984, Spannaus in 1990, and M. Warner in 1996.

³ In Texas, Tower (R) defeated Krueger (D) in 1978; Gramm (R) defeated Doggett in 1984, Parmer in 1990, and Morales in 1996.

⁴ In Kentucky, McConnell (R) defeated Huddleston (D) in 1984, Sloane in 1990, and Beshear in 1996 (Class 2 election).

⁵ In Florida, Mack (R) defeated MacKay (D) in 1988, and Rodham in 1994; McCollum (R) defeated Nelson (D) in 2000.

⁶ In Wyoming, Wallop (R) defeated Vinich (D) in 1988, and Thomas (R) defeated Sullivan in 1994.

⁷ In Georgia, Coverdell (R) defeated Fowler (D) in 1992, and Coles in 1998. After Coverdell's death in 2000, he was succeeded by Miller (D).

⁸ In Kentucky, Bunning (R) defeated Baesler (D) in 1998 (Class 3 election).

Table 3: Logistic Regression Predicting 1996 Voter Turnout and 1996 and 1998 Party Preference

		1996 TU	1996-1998 P	REFERENCE		
Voting Predictors	Model 1	Model 2	Model 3	Model 4	Clinton (D)	Ventura (I)
Criminal Sanction						
Any arrest	681**	264				
	(.217)	(.252)				
Property arrest			323	.148	242	346
			(.326)	(.353)	(.488)	(.597)
Drug/alcohol arrest			341	171	1.274*	1.599*
-			(.342)	(.380)	(.633)	(.789)
Violent arrest			-1.246*	851	758	.946
			(.501)	(.541)	(.860)	(1.150)
Other arrest			065	.145	.582	.198
			(.372)	(.397)	(.589)	(.771)
Voting Predictors			,	,	,	,
Nonwhite (vs. white)		663**		628**	1.216*	792
,		(.258)		(.261)	(.517)	(.422)
Female		.066		.089	1.231*	332
		(.216)		(.215)	(.266)	(.281)
Years education		.415**		.414**	.117	536**
		(.063)		(.063)	(.085)	(.102)
Income (in thousands)		.036**		.036**	004	.001
,		(.012)		(.012)	(.014)	(.016)
Full-time employment		257		268	390	592
1 3		(.240)		(.240)	(.313)	(.342)
Married		.088		.018	.130	.076
		(.224)		(.223)	(.293)	(.301)
Constant	.928**	-5.429**	.879**	-5.452**	-1.228	8.778**
	(.107)	(.925)	(.103)	(.923)	(1.281)	(1.554)
Number of Cases	548	548	550	550	354	285
-2 Log Likelihood	673.8**	599.4**	676.1**	603.4**	373.6**	368.7**

Note: Standard errors in parentheses *p < .05 **p < .01

Table 4. Applying Contemporary Disenfranchisement Rates to the 2000 and 1960 Presidential Elections.

A. 2000 Election: What if Felons had been Allowed to Vote?

	Actual	Total	Est.	Est.	Net	Counter-
	(R)	Disenfran-	Turnout	Percent	Democrat	factual (D)
Unit	Margin	chised	Rate	Democrat	Votes Lost	Margin
U.S. Total	-539,947	4,695,729	29.7%	68.9%	527,171	1,067,118
Florida	537	827,207	27.2%	68.9%	85,050	84,513
50% Lower Turnout			13.6%	68.9%	42,525	41,988
FL Ex-Felons Only		613,514	27.2%	68.9%	63,079	62,542
50% Lower Turnout			13.6%	68.9%	31,540	31,003

B. 1960 Election: What if Felons were Disenfranchised in 1960 at 2000 Rates?

	Actual	Actual	Counterfactual	Est.	Est.	Net	Counter-
	(D)	Disenfran-	Disenfranchised	Turnout	Percen	Democrat	factual (R)
	Margin	chised		Rate	t Dem.	Votes Lost	Margin
U.S. Total	118,550	1,378,156	2,502,211	40%	75%	224,811	106,261
50% Lower Turnout				20%	75%	112,405	(6,145)

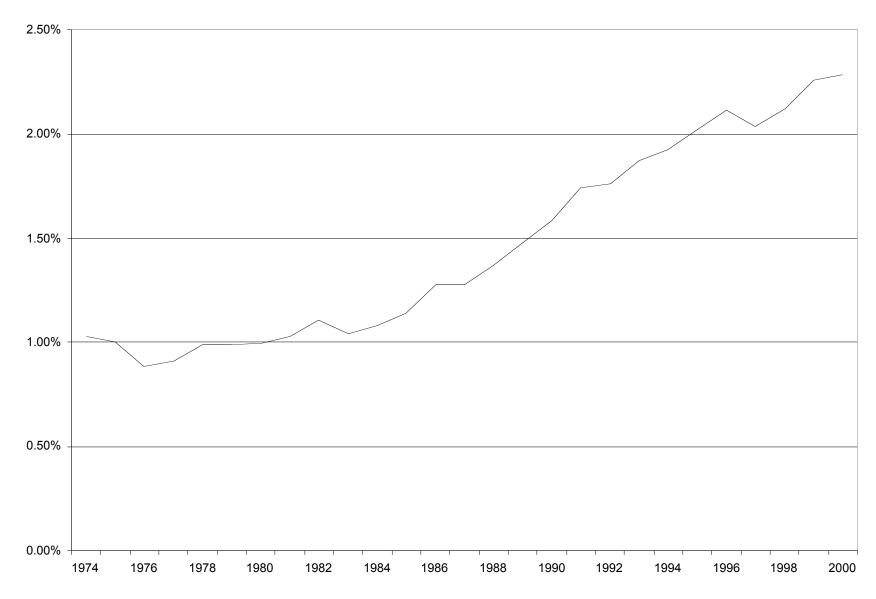
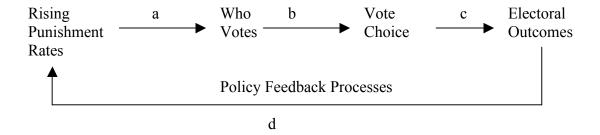


Figure 1. Felon Disenfranchisement as a Percentage of the U.S. Voting Age Population, 1974-2000

Figure 2. What is the Impact of Felon Disenfranchisement on American Politics?



Percentage of States with Any Felon Disfranchisement Provisi

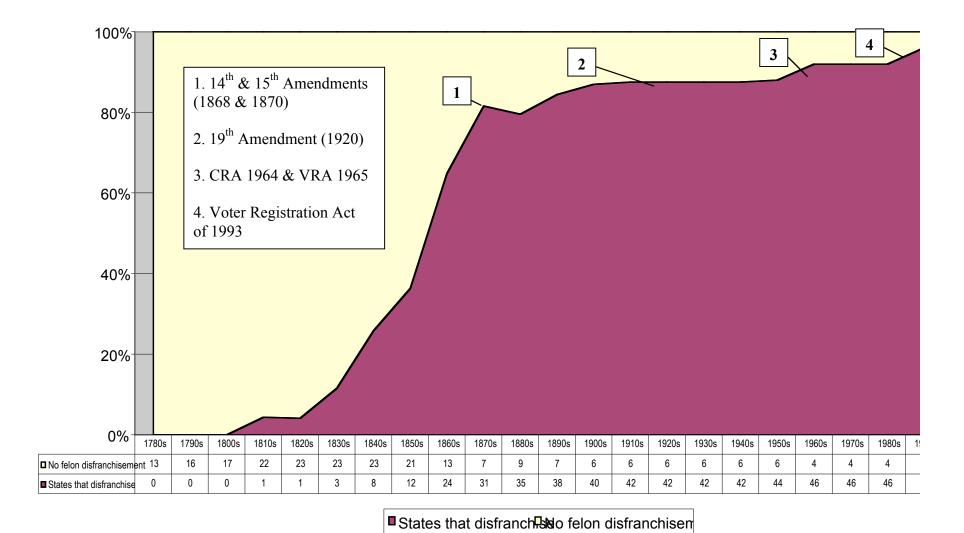


Figure 3. The Prevalence of Felon Disenfranchisement in the United States, 1788-2000.

APPENDIX TABLE 1: Estimates of Disenfranchised Felons by State 12/31/00.

Alabama	State	Prisoners	Parolees	Felony	Jail	Estimated	Total	Voting Age	Disf.
Alabama	State	1 Historici s	1 ai dices				1 Otal		
Alaska 2,128 507 4,543 212 7,390 430,000 1,72% Arizona 26,510 3,474 50,897 1,033 58,936 1,140,00 2,619,00 2,61% California 163,001 117,647 7,714 28,862 24,873,000 1,16% Connecticut 13,155 1,868 29,641 520 45,184 2,499,000 1,81% Delaware 3,937 579 10,808 298 14,384 30,006 882,000 5,18% Dist, Col 7,456 143 7,599 411,000 7,038 Florida 7,1233 6,046 13,186 5,228 613,514 82,7207 11,774,000 7,038 Georgia 44,232 21,556 217,038 3,451 286,277 583,000 4,86% Hawaii 3,550 1,443 8,774 321 1,604 92,100 1,74% Illinois 45,281 1,443 8,734 3,21 4,604	Alahama	26 225	5 404				212 650		
Arizona 26,510 3,474 50,897 1,053 58,936 140,870 3,625,000 2,616 California 163,001 117,647 7,714 288,362 24,873,000 1,16% Colorado 16,833 5,500 967 23,300 3,067,000 0.76% Comecticut 13,155 1,868 29,641 520 43,184 24,990,000 1,11% Delaware 3,937 579 10,808 298 14,384 30,006 582,000 5,16% Florida 71,233 6,046 131,186 5,228 613,514 827,207 11,774,000 7,03% Georgia 44,232 21,556 217,038 3,451 286,277 5893,000 4,06% Hawaii 3,553 1,50 3,703 909,000 0.41% Idaho 5,266 1,443 8,774 321 16,064 921,000 4,96% Illinois 4,228 1,333 3,025 10,063 21,656 <t< td=""><td></td><td></td><td></td><td></td><td></td><td>140,030</td><td></td><td></td><td></td></t<>						140,030			
Arkansas						58 036			
California 103,001 117,647 7,714 288,362 24,873,000 0,76% Colorado 16,833 5,500 967 23,300 3,067,000 0,76% Connecticut 13,155 1,868 29,641 520 45,184 2,499,000 1,81% Dist Col 7,456 70 10,808 298 14,384 30,006 582,000 5,16% Brorida 71,233 6,046 131,186 5,228 613,514 827,207 11,740,000 7,03% Georgia 44,232 21,558 217,038 3,451 827,207 11,774,000 7,03% Hawaii 3,553 150 3,703 909,000 0,41% Idaho 5,526 1,443 8,774 321 16,64 4291,000 1,74% Ilminois 45,281 1,741 46,992 8,983,000 0,52% Kansas 3,344 3,829 426 12,599 1,983,000 0,64% Kanisas 1,44					1,033	36,930			
Colorado 16,833 5,500 967 23,300 3,067,000 0.76% Connecticut 13,155 1,868 29,641 520 43,184 2,499,000 1.81% Delaware 3,937 579 10,808 298 14,384 30,006 582,000 5.16% Dist Col 7,456 17,433 6,046 131,186 5,228 613,514 827,207 17,470,000 7.03% Georgia 44,232 21,556 217,038 3,451 286,277 5,893,000 4.8% Idwaii 3,553 150 3,703 909,000 0.48% Idwaii 45,281 1,711 46,992 8,983,000 0.52% Illinois 45,281 1,333 80,257 106,61 21,569,000 0.48% Kentucky 14,919 4,909 17,464 1,010 109,132 14,743 2,93,000 0.4% Kentucky 14,919 4,909 17,464 1,010 109,132 14,743				29,046	7 714				
Connecticut 13,155 1,868 29,641 520 45,184 2,499,000 1,81% Delaware 3,937 579 10,808 298 14,384 30,006 582,000 5,16% Florida 71,233 6,046 131,186 5,228 613,514 827,207 11,774,000 7,03% Georgia 44,232 21,556 217,038 3,451 826,277 5,893,000 0,43% Hawaii 3,553 150 3,703 909,000 0,41% Idaho 5,526 1,443 8,774 321 16,664 921,000 1,74% Ildian 20,125 1,333 80,257 100,631 2,61600 0,448 Ilowa 7,955 2,763 9,326 330 80,257 100,631 2,993,000 0,49% Kansas 8,344 3,829 426 12,599 1,983,000 0,64% Kentucky 14,919 4,909 17,464 1,010 109,122 147,434									
Delaware				20.641					
Dist. Col 7,456 40,466 131,186 5,228 613,514 827,207 11,774,000 7,03% Georgia 44,232 21,556 217,038 3,451 286,277 5,893,000 0.48% Hawaii 3,553 150 3,703 909,000 0.14% Idaho 5,526 1,443 8,774 321 16,064 921,000 1,74% Illinois 45,281 1,333 21,458 8,448,000 0.48% Ilowa 7,955 2,763 9,326 330 80,257 100,631 2,165,000 0.45% Kansas 8,344 3,829 426 12,599 1,983,000 0.04% Kentucky 14,919 4,909 17,464 1,010 109,132 147,434 325,500 1,16% Mair 15,007 2,637 37,684 32,55,000 1,16% Maryland 23,538 14,143 22,563 1,115 78,206 139,565 3,025,000 3,56% <t< td=""><td></td><td></td><td></td><td></td><td></td><td>14204</td><td></td><td></td><td></td></t<>						14204			
Florida			3/9	10,808		14,384			
Georgia 44,232 21,556 217,038 3,451 286,277 5,893,000 4,86% Hawaii 3,553 1,443 8,774 321 16,064 921,000 1,74% Illinois 45,281 1,711 46,992 8,983,000 0,52% Ilmiana 20,125 1,333 21,458 4,448,000 0,48% Iowa 7,955 2,763 9,326 330 80,257 100,631 2,165,000 4,65% Kansas 8,344 3,829 426 12,599 1,983,000 0,63% Kentucky 14,919 4,909 17,464 1,010 109,132 147,434 2,993,000 0,63% Maine 2,637 37,684 3,255,000 1,10% Maryland 23,538 14,143 22,563 1,115 78,206 139,565 3,925,000 0,0% Mississina 47,718 1,600 49,318 7,358,000 0,67% Mismesota 6,238 3,072 3			6.046	121 106		612.514			
Hawaii 3.553 1.443 8,774 321 16,064 921,000 1.74% Illinois 45,281 1,711 46,992 8,983,000 0.52% Indiana 20,125 1,333 21,458 4,448,000 0.48% Iowa 7,955 2,763 9,326 330 80,257 100,631 2,165,000 0.65% Kansas 8,344 3,829 426 12,599 1,983,000 0.64% Kentucky 14,919 4,909 17,464 1,010 109,132 147,434 2,993,000 0.49% Maine 2,637 78,206 139,565 3,255,000 1.66% Maryland 23,538 14,143 22,563 1,115 78,206 139,565 3,925,000 3.56% Missouri 4,718 4,648 523 41,477 3,44,479 0.00% Minesotra 2,231 1,566 15,118 986 82,002 119,943 2,047,000 5.86% Missouri<						613,514			
Idaho 5,526 1,443 8,774 321 16,064 921,000 1,74% Illinois 45,281 1,711 46,992 8,983,000 0.52% Indiana 20,125 1,333 21,458 4,448,000 0.48% Iowa 7,955 2,763 9,326 330 80,257 100,631 2,165,000 4,65% Kansas 8,344 3,829 426 12,599 1,983,000 0.64% Kentucky 14,919 4,909 17,464 1,010 109,132 147,434 2,993,000 493% Louisiana 35,047 2,637 37,684 3,255,000 1,16% Maine - 6 98,800 0.00% Mass. - 1,600 49,318 7,358,00 0.07% Mississipin 47,718 1,600 49,318 7,350 0.074,790 0.00% Mississippi 20,241 1,596 15,118 986 82,002 119,943 2,047,000	_		21,556	217,038					
Illinois			1 112	0.554					
Indiana 20,125			1,443	8,774					
Iowa 7,955 2,763 9,326 330 80,257 100,631 2,165,000 4,65% Kansas 8,344 3,829 426 12,599 1,983,000 0,64% Kentucky 14,919 4,909 17,464 1,010 109,132 147,434 2,930,000 4,93% Louisiana 35,047 2,637 78,206 139,565 3,925,000 3,56% Maryland 23,538 14,143 22,563 1,115 78,206 139,565 3,925,000 3,56% Mass: 0 4,749,000 0,00% Michigan 47,718 1,600 49,318 7,358,000 0,67% Missouri 27,323 12,357 42,607 725 83,012 4,105,000 5,86% Missouri 27,323 12,357 42,607 725 83,012 4,074,000 5,86% Missouri 27,323 4,2667 725 83,012 4,105,000 2,28% Neb Taska 3,895									
Kansas 8,344 3,829 426 12,599 1,983,000 0.64% Kentucky 14,919 4,909 17,464 1,010 109,132 147,434 2,993,000 4,93% Louisiana 35,047 2,637 37,684 3,255,000 1,16% Maine 0 968,000 0,00% Maryland 23,538 14,143 22,563 1,115 78,206 139,565 3,925,000 3,56% Miss 47,718 1,600 49,318 7,358,000 0,67% Minnesota 6,238 3,072 31,644 523 41,477 3,547,000 1,78% Mississippi 20,241 1,596 15,118 986 82,002 119,943 2,047,000 5,86% Missouri 27,323 12,357 42,607 725 83,012 4,105,000 2,02% Montana 3,105 4,828 231 44,01 53,428 1,311,30 1,32 668,000 0,47% New Jac									
Kentucky 14,919 4,909 17,464 1,010 109,132 147,434 2,993,000 4,93% Louisiana 35,047 2,637 37,884 3,255,000 1.16% Maryland 23,538 14,143 22,563 1,115 78,206 139,565 3,925,000 3,56% Mass. 0 4,749,000 0,00% Michigan 47,718 1,600 49,318 7,358,000 0,67% Minnesota 6,238 3,072 31,644 523 41,477 3,547,000 1.17% Mississippi 20,241 1,596 15,118 986 82,002 119,943 2,047,000 5.86% Missouri 27,323 12,357 42,607 725 83,012 4,105,000 2,02% Montana 3,105 473 4,828 231 44,001 53,428 1,234,000 4,33% Nevardac 10,012 4,056 8,410 517 43,395 66,390 1,394 4,78%			,	9,326		80,257			
Louisiana Maine 2,637 2,637 37,684 3,255,000 1.16% Maine 0 968,000 0.0% Maryland 23,538 14,143 22,563 1,115 78,206 139,565 3,925,000 3.56% Mass. 0 4,749,000 0.00% Michigan 47,718 1,600 49,318 7,358,000 0.67% Minnesota 6,238 3,072 31,644 523 41,477 3,547,000 1.17% Mississippi 20,241 1,596 15,118 986 82,002 119,943 2,047,000 5.86% Missouri 27,323 12,357 42,607 725 83,012 4,105,000 2.02% Missouri 27,323 12,357 42,607 725 83,012 4,105,000 2.02% Montana 3,105 160 3,265 668,000 0.49% Mehraska 3,895 473 4,828 231 44,001 53,428 1,234,000 4,33% Nevada 10,012 4,056 8,410 517 43,395 66,390 1,390,000 4,78% New Hamp. 2,257 159 2,416 911,000 2.27% New Jersey 29,784 14,899 96,831 1,592 143,106 6,245,000 2.29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 6,21% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 0.24% Ohio 45,833 1,628 47,461 8,433,000 0.56% Oregon 10,630 45,833 1,584 132 1,584 1,525 1,586 1,586 1,586 1,586 1,586 1,586 1,586 1,586 1,586 1,586 1,586 1,586 1,586									
Maine Goldward 968,000 0.00% Maryland 23,538 14,143 22,563 1,115 78,206 139,565 3,925,000 3,56% Mass. 0 47,718 1,600 49,318 7,358,000 0.67% Minnesota 6,238 3,072 31,644 523 41,477 3,547,000 1.17% Mississippi 20,241 1,596 15,118 986 82,002 119,943 2,047,000 5.86% Missouri 27,323 12,357 42,607 725 83,012 4,105,000 2.02% Montana 3,105 42,607 725 83,012 4,105,000 2.02% Mevada 10,012 4,056 8,410 517 43,395 66,390 4,33% New Hamp. 2,257 159 2,416 911,000 2.27% New Horico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 6.21% New York 70,198	•		4,909	17,464		109,132			
Maryland 23,538 14,143 22,563 1,115 78,206 139,565 3,925,000 3.56% Mass. 0 4,749,000 0,00% Michigan 47,718 1,600 49,318 7,358,000 0.67% Minnesota 6,238 3,072 31,644 523 41,477 3,547,000 1.17% Missouri 27,323 12,357 42,607 725 83,012 4,105,000 2.02% Montana 3,105 160 3,265 668,000 0.49% Nebraska 3,895 473 4,828 231 44,001 53,428 1,234,000 4,33% New Hamp. 2,257 159 2,416 911,000 0.27% New Jersey 29,784 14,899 96,831 1,592 143,106 6,245,000 2,29% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334		35,047			2,637				
Mass. Michigan 47,718 1,600 49,318 7,358,000 0.60% of 0.6	Maine								
Michigan 47,718 (6.28) 1,600 (7.28) 49,318 (7.358,000 (7.28) 0.67% (7.28) Minnesota 6,238 (3.072) 31,644 (523) 41,477 (3.547,000 (1.17%) 1.17% (7.28) Mississippi 20,241 (1.596 (15,118) 986 (82,002) 119,943 (2.047,000) 5.86% (7.28) Missouri 27,323 (12,357) 42,607 (725) 83,012 (4.105,000) 2.02% (7.28) Montana 3,105 (7.28) 42,607 (725) 83,012 (4.105,000) 2.02% (7.28) Mortana 3,105 (7.28) 473 (4.828) 231 (44,001) 33,428 (1.234,000) 4.33% (7.28) Nevada 10,012 (4.056) 8,410 (517) 43,395 (6.390) 1,390,000 (4.78%) 4.78% (7.28%) New Hamp. 2,257 (7.279) 159 (7.279) 2,416 (911,000) (2.29%) 2.29% (7.28%) 4.331,000 (6.21%) 6.245,000 (2.29%) 2.29% (7.28%) New York (7.198 (7.28%) 3,217 (7.279) (3.44 (7.28%) 4.31,273 (3.3805,000) (6.21%) 6.245,000 (2.29%) 0.24% (7.28%) 0.24% (7.28%) 0.13,273 (3.3805,000) (6.21%) 0.24% (7.28%) 0.13,273 (3.38,000) (7.5%) 0.24% (7.28%) 0.13,273 (3.38,000) (7.5%) 0.24% (7.28%) 0.24% (7.28%) <td>Maryland</td> <td>23,538</td> <td>14,143</td> <td>22,563</td> <td>1,115</td> <td>78,206</td> <td>139,565</td> <td>3,925,000</td> <td>3.56%</td>	Maryland	23,538	14,143	22,563	1,115	78,206	139,565	3,925,000	3.56%
Minnesota 6,238 3,072 31,644 523 41,477 3,547,000 1.17% Mississippi 20,241 1,596 15,118 986 82,002 119,943 2,047,000 5.86% Missouri 27,323 12,357 42,667 725 83,012 4,105,000 2.02% Montana 3,105 160 3,265 668,000 0.49% Nebraska 3,895 473 4,828 231 44,001 53,428 1,234,000 4,339 New Hamp. 2,257 159 2,416 911,000 0,27% New Jersey 29,784 14,899 96,831 1,592 143,106 6,245,000 2,29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 0,25% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,979,000 1,22% North Dakota 1,076 67 1,143 47,601 8,433,000	Mass.						0	4,749,000	0.00%
Mississippi 20,241 1,596 15,118 986 82,002 119,943 2,047,000 5.86% Missouri 27,323 12,357 42,607 725 83,012 4,105,000 2.02% Montana 3,105 160 3,265 668,000 0.49% Nebraska 3,895 473 4,828 231 44,001 53,428 1,234,000 4,33% Nevada 10,012 4,056 8,410 517 43,395 66,390 1,390,000 4.78% New Hamp. 2,257 159 2,416 911,000 0.27% New Jersey 29,784 14,899 96,831 1,592 143,106 62,45,000 2.29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 6.21% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% North Dakota 1,076 67 1,143 477,000 0.26%	Michigan	47,718			1,600		49,318	7,358,000	0.67%
Missouri 27,323 12,357 42,607 725 83,012 4,105,000 2.02% Montana 3,105 6 160 3,265 668,000 0.49% Nebraska 3,895 473 4,828 231 44,001 53,428 1,234,000 4,33% New Acada 10,012 4,056 8,410 517 43,395 66,390 1,390,000 4,78% New Hamp. 2,257 159 2,416 911,000 0.27% New Jersey 29,784 14,899 96,831 1,592 143,106 62,45,000 2,29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 6.21% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1,22% North Dakota 1,076 67 1,143 47,000 0,24%	Minnesota	6,238	3,072	31,644	523		41,477	3,547,000	1.17%
Montana 3,105 473 4,828 231 44,001 53,428 1,234,000 4,33% Nevada 10,012 4,056 8,410 517 43,395 66,390 1,390,000 4.78% New Hamp. 2,257 159 2,416 911,000 0.27% New Jersey 29,784 14,899 96,831 1,592 143,106 6,245,000 2.29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 6.21% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1.22% North Dakota 1,076 67 1,143 477,000 0.24% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pen	Mississippi	20,241	1,596	15,118	986	82,002	119,943	2,047,000	5.86%
Nebraska 3,895 473 4,828 231 44,001 53,428 1,234,000 4,33% Nevada 10,012 4,056 8,410 517 43,395 66,390 1,390,000 4,78% New Hamp. 2,257 159 2,416 911,000 0.27% New Jersey 29,784 14,899 96,831 1,592 143,106 6,245,000 2.29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 0.21% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1.22% North Dakota 1,076 67 1,143 477,000 0.24% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pe	Missouri	27,323	12,357	42,607	725		83,012	4,105,000	2.02%
Nevada 10,012 4,056 8,410 517 43,395 66,390 1,390,000 4.78% New Hamp. 2,257 159 2,416 911,000 0.27% New Jersey 29,784 14,899 96,831 1,592 143,106 6,245,000 2.29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 62.1% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1.22% North Dakota 1,076 67 1,143 477,000 0.24% Ohio 45,833 683 698 52,089 2,531,000 2.06% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Pennsylvania 36,847 9,155,000 0.40% 11,307 2,530,000 0.45% Ro. Carolina	Montana	3,105			160		3,265	668,000	0.49%
New Hamp. 2,257 159 2,416 911,000 0.27% New Jersey 29,784 14,899 96,831 1,592 143,106 6,245,000 2.29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 6.21% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1.22% North Dakota 1,076 67 1,143 477,000 0.24% Ohio 45,833 1,628 47,461 8,433,000 0.56% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pennsylvania 36,847 36,847 91,550,000 0.40% Rhode Island 1,966 353 15,844 132 18,295 753,000	Nebraska	3,895	473	4,828	231	44,001	53,428	1,234,000	4.33%
New Hamp. 2,257 159 2,416 911,000 0.27% New Jersey 29,784 14,899 96,831 1,592 143,106 6,245,000 2.29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 6.21% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1.22% North Dakota 1,076 67 1,143 477,000 0.24% Ohio 45,833 1,628 47,461 8,433,000 0.56% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Rhode Island 1,966 353 15,844 132 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323	Nevada	10,012	4,056	8,410	517	43,395	66,390	1,390,000	4.78%
New Jersey 29,784 14,899 96,831 1,592 143,106 6,245,000 2.29% New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 6.21% New York 70,198 57,858 3,217 131,273 13,805,000 0,95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1.22% Ohio 45,833 67 67 1,143 477,000 0.24% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pennsylvania 36,847 36,847 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 111 2,727 542,000 0.50% Texas 157,997 111,719	New Hamp.	2,257			159		2,416	911,000	0.27%
New Mexico 5,342 1,670 7,279 544 63,565 78,400 1,263,000 6.21% New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1.22% North Dakota 1,076 67 1,143 477,000 0.24% Ohio 45,833 1,628 47,461 8,433,000 0.56% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pennsylvania 36,847 9,155,000 0.40% Rhode Island 1,966 353 15,844 132 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 8,094 30,235 1,934 28,720 <td></td> <td>29,784</td> <td>14,899</td> <td>96,831</td> <td>1,592</td> <td></td> <td>143,106</td> <td>6,245,000</td> <td>2.29%</td>		29,784	14,899	96,831	1,592		143,106	6,245,000	2.29%
New York 70,198 57,858 3,217 131,273 13,805,000 0.95% No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1.22% North Dakota 1,076 67 1,143 477,000 0.24% Ohio 45,833 1,628 47,461 8,433,000 0.56% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pennsylvania 36,847 36,847 36,847 9,155,000 0.40% Rhode Island 1,966 353 15,844 132 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 111 2,727 542,000 0.50% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000	•					63,565			
No. Carolina 31,266 3,352 34,701 1,334 70,653 5,797,000 1.22% North Dakota 1,076 67 1,143 477,000 0.24% Ohio 45,833 1,628 47,461 8,433,000 0.56% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pennsylvania 36,847 36,847 9,155,000 0.40% Rhode Island 1,966 353 15,844 132 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 111 2,727 542,000 0.50% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Vermont 0 460,000 0.00% 0.61% 460,000 0.61% <				Ź	3,217	,			
North Dakota 1,076 67 1,143 477,000 0,24% Ohio 45,833 1,628 47,461 8,433,000 0.56% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pennsylvania 36,847 36,847 9,155,000 0.40% Rhode Island 1,966 353 15,844 132 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 111 2,727 542,000 0.50% Tennessee 22,166 8,094 30,235 1,934 28,720 91,149 4,221,000 2.16% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Utah 5,630 3,266 8,896 1,465,000 0.61% </td <td></td> <td></td> <td></td> <td>34,701</td> <td></td> <td></td> <td>,</td> <td></td> <td></td>				34,701			,		
Ohio 45,833 1,628 47,461 8,433,000 0.56% Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pennsylvania 36,847 9,155,000 0.40% Rhode Island 1,966 353 15,844 132 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 111 2,727 542,000 0.50% Tennessee 22,166 8,094 30,235 1,934 28,720 91,149 4,221,000 2.16% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Utah 5,630 3,266 1,847 243,902 310,661 5,263,000 5.90% Washington 14,915 160 109,956 1,078			- ,	- ,				, ,	
Oklahoma 23,181 1,825 26,385 698 52,089 2,531,000 2.06% Oregon 10,630 677 11,307 2,530,000 0.45% Pennsylvania 36,847 9,155,000 0.40% Rhode Island 1,966 353 15,844 132 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 111 2,727 542,000 0.50% Tennessee 22,166 8,094 30,235 1,934 28,720 91,149 4,221,000 2.16% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Utah 5,630 3,266 8,896 1,465,000 0.61% Vermont 0 460,000 0.00% Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% <									
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Pennsylvania 36,847 9,155,000 0.40% Rhode Island 1,966 353 15,844 132 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 111 2,727 542,000 0.50% Tennessee 22,166 8,094 30,235 1,934 28,720 91,149 4,221,000 2.16% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Utah 5,630 3,266 8,896 1,465,000 0.61% Vermont 0 460,000 0.00% Virginia 30,168 5,148 29,596 1,847 243,902 310,661 5,263,000 5.90% Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% West Virginia 3,856 1,112 3,635 272			1,025	20,505					
Rhode Island 1,966 353 15,844 132 18,295 753,000 2.43% So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 111 2,727 542,000 0.50% Tennessee 22,166 8,094 30,235 1,934 28,720 91,149 4,221,000 2.16% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Utah 5,630 3,266 8,896 1,465,000 0.61% Vermont 0 460,000 0.00% Virginia 30,168 5,148 29,596 1,847 243,902 310,661 5,263,000 5.90% Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% West Virginia 3,856 1,112 3,635 272 8,875 1,416,000 0.63% Wisconsin					077				
So. Carolina 21,778 4,240 25,323 869 52,210 2,977,000 1.75% South Dakota 2,616 111 2,727 542,000 0.50% Tennessee 22,166 8,094 30,235 1,934 28,720 91,149 4,221,000 2.16% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Utah 5,630 3,266 8,896 1,465,000 0.61% Vermont 0 460,000 0.00% Virginia 30,168 5,148 29,596 1,847 243,902 310,661 5,263,000 5.90% Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% West Virginia 3,856 1,112 3,635 272 8,875 1,416,000 0.63% Wisconsin 20,612 9,430 22,715 1,268 54,025 3,930,000 1.37% Wyoming			353	15 844	132				
South Dakota 2,616 111 2,727 542,000 0.50% Tennessee 22,166 8,094 30,235 1,934 28,720 91,149 4,221,000 2.16% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Utah 5,630 3,266 8,896 1,465,000 0.61% Vermont 0 460,000 0.00% Virginia 30,168 5,148 29,596 1,847 243,902 310,661 5,263,000 5.90% Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% West Virginia 3,856 1,112 3,635 272 8,875 1,416,000 0.63% Wisconsin 20,612 9,430 22,715 1,268 54,025 3,930,000 1.37% Wyoming 1,680 514 2,760 99 12,797 17,850 358,000 4.99%									
Tennessee 22,166 8,094 30,235 1,934 28,720 91,149 4,221,000 2.16% Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Utah 5,630 3,266 8,896 1,465,000 0.61% Vermont 0 460,000 0.00% Virginia 30,168 5,148 29,596 1,847 243,902 310,661 5,263,000 5.90% Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% West Virginia 3,856 1,112 3,635 272 8,875 1,416,000 0.63% Wisconsin 20,612 9,430 22,715 1,268 54,025 3,930,000 1.37% Wyoming 1,680 514 2,760 99 12,797 17,850 358,000 4.99%		21,776	7,270	23,323					
Texas 157,997 111,719 250,642 5,609 525,967 14,850,000 3.54% Utah 5,630 3,266 8,896 1,465,000 0.61% Vermont 0 460,000 0.00% Virginia 30,168 5,148 29,596 1,847 243,902 310,661 5,263,000 5.90% Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% West Virginia 3,856 1,112 3,635 272 8,875 1,416,000 0.63% Wisconsin 20,612 9,430 22,715 1,268 54,025 3,930,000 1.37% Wyoming 1,680 514 2,760 99 12,797 17,850 358,000 4.99%			9 004	20.225		29 720			
Utah 5,630 3,266 8,896 1,465,000 0.61% Vermont 0 460,000 0.00% Virginia 30,168 5,148 29,596 1,847 243,902 310,661 5,263,000 5.90% Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% West Virginia 3,856 1,112 3,635 272 8,875 1,416,000 0.63% Wisconsin 20,612 9,430 22,715 1,268 54,025 3,930,000 1.37% Wyoming 1,680 514 2,760 99 12,797 17,850 358,000 4.99%				,		20,720			
Vermont 0 460,000 (0.00%) Virginia 30,168 (5,148) (29,596) (1,847) (243,902) (310,661) (5,263,000) (5.90%) Washington 14,915 (160) (109,956) (1,078) (1,078) (32,856) (158,965) (4,368,000) (3.64%) West Virginia 3,856 (1,112) (3,635) (272) (272) (8,875) (1,416,000) (0.63%) Wisconsin 20,612 (9,430) (22,715) (1,268) (1,268) (54,025) (3,930,000) (1,37%) Wyoming 1,680 (514) (2,760) (99) (12,797) (17,850) (358,000) (4.99%)				230,042	3,009				
Virginia 30,168 5,148 29,596 1,847 243,902 310,661 5,263,000 5.90% Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% West Virginia 3,856 1,112 3,635 272 8,875 1,416,000 0.63% Wisconsin 20,612 9,430 22,715 1,268 54,025 3,930,000 1.37% Wyoming 1,680 514 2,760 99 12,797 17,850 358,000 4.99%		3,030	3,200						
Washington 14,915 160 109,956 1,078 32,856 158,965 4,368,000 3.64% West Virginia 3,856 1,112 3,635 272 8,875 1,416,000 0.63% Wisconsin 20,612 9,430 22,715 1,268 54,025 3,930,000 1.37% Wyoming 1,680 514 2,760 99 12,797 17,850 358,000 4.99%		20.169	£ 140	20.506	1 0 4 7	242.002			
West Virginia 3,856 1,112 3,635 272 8,875 1,416,000 0.63% Wisconsin 20,612 9,430 22,715 1,268 54,025 3,930,000 1.37% Wyoming 1,680 514 2,760 99 12,797 17,850 358,000 4.99%							,		
Wisconsin 20,612 9,430 22,715 1,268 54,025 3,930,000 1.37% Wyoming 1,680 514 2,760 99 12,797 17,850 358,000 4.99%						32,856			
Wyoming 1,680 514 2,760 99 12,797 17,850 358,000 4.99%									
						10 =0=			
Total 1,209,243 444,405 1,320,684 57,710 1,654,497 4,686,539 205,814,000 2.28%									
	Total	1,209,243	444,405	1,320,684	57,710	1,654,497	4,686,539	205,814,000	2.28%

APPENDIX TABLE 2: Estimates of Disenfranchised African American Felons by State 12/31/00.

APPENDIX TABLE 2: Estimates of Disenfranchised African American Felons by State 12/31/00. State Black Black Black Black Est. Black Voting Age												
State						T . 1		Disf.				
	Prison	Parole	Probation	Jail	Ex Felons	Total	Population	Rate				
Alabama	17,230	2,674	13,248	671	77,932	111,755	800,000	13.97%				
Alaska	317	53	585	10		966	17,000	5.68%				
Arizona	4,016	543	4,347	143	8,651	17,700	137,000	12.92%				
Arkansas	6,595	4,715	10,376			21,686	276,000	7.86%				
California	80,490	31,457		2,697		114,644	1,853,000	6.19%				
Colorado	4,224	1,639		199		6,063	132,000	4.59%				
Connecticut	8,302	1,175	8,689	250		18,417	221,000	8.33%				
Delaware	2,524	303	5,069		7,162	15,058	108,000	13.94%				
Dist. Col	7,382			131		7,513	230,000	3.27%				
Florida	39,427	3,472	43,305	2,774	167,413	256,392	1,600,000	16.02%				
Georgia	29,583	14,267	115,711	2,124		161,685	1,577,000	10.25%				
Hawaii	201			6		208	27,000	0.77%				
Idaho	105	28	141	6		280	7,000	4.00%				
Illinois	32,780			1,116		33,895	1,249,000	2.71%				
Indiana	8,664			634		9,297	353,000	2.63%				
Iowa	2,028	411	1,019	62	7,671	11,192	45,000	24.87%				
Kansas	3,218	1,359		117		4,694	112,000	4.19%				
Kentucky	5,718	1,377	3,916	312	24,632	35,955	207,000	17.37%				
Louisiana	26,820	ŕ	ŕ	1,870	ŕ	28,690	956,000	3.00%				
Maine							7,000	0.00%				
Maryland	18,228	10,662	13,105	736	42,519	85,251	1,058,000	8.06%				
Mass.	Ź	,	Ź		Ź	,	270,000	0.00%				
Michigan	27,230			572		27,802	977,000	2.85%				
Minnesota	2,309	1,841	4,587	128		8,865	106,000	8.36%				
Mississippi	15,145	1,130	9,099	698	50,035	76,106	675,000	11.27%				
Missouri	12,489	4,964	12,719	300	Ź	30,471	425,000	7.17%				
Montana	44	,	Ź	4		48	4,000	1.21%				
Nebraska	1,155	116	758	47	7,164	9,240	49,000	18.86%				
Nevada	3,118	1,331	1,853	154	11,514	17,970	105,000	17.11%				
New Hamp.	125	,	,	12	Ź	138	9,000	1.53%				
New Jersey	21,301	8,977	47,666	975		78,920	856,000	9.22%				
New Mexico	621	199	515	43	7,750	9,128	37,000	24.67%				
New York	38,849	43,638		1,749	,	84,236	2,309,000	3.65%				
No. Carolina	20,480	2,114	17,448	868		40,910	1,173,000	3.49%				
North Dakota	27	,	Ź	2		29	4,000	0.72%				
Ohio	24,829			720		25,549	895,000	2.85%				
Oklahoma	8,336	614	6,108	225		15,283	185,000	8.26%				
Oregon	1,506		.,	74		1,580	51,000	3.10%				
Pennsylvania	23,104					23,104	820,000	2.82%				
Rhode Island	685	100	3,598	35		4,419	36,000	12.27%				
So. Carolina	15,262	2,949	13,950	596		32,756	816,000	4.01%				
South Dakota	116	_,, .,	,	3		119	5,000	2.37%				
Tennessee	11,277	4,605	12,806	1,125	11,946	41,759	635,000	6.58%				
Texas	71,915	44,282	46,546	2,130	11,5 .0	164,873	1,800,000	9.16%				
Utah	432	244	10,5 10	2,130		676	16,000	4.23%				
Vermont	132	211				070	4,000	0.00%				
Virginia	20,234	3,323	15,085	1,180	121,737	161,559	1,005,000	16.08%				
Washington	3,376	23	14,647	205	3,824	22,075	154,000	14.33%				
West Virginia	615	218	316	39	3,024	1,188	45,000	2.64%				
Wisconsin	9,940	4,476	5,920	469		20,805	193,000	10.78%				
Wyoming	101	4,470	3,920	2	358	20,803	4,000	14.18%				
Total	632,474	199,301	433,216	26,215	550,308	1,841,515	24,635,000	7.48%				

Appendix Table 3: Logistic Regression Coefficients for Turnout and Democratic Senate and Presidential Vote, 1972 – 2000

TURNOUT ANALYSIS (voted = 1)

	<u> 1972</u>	<u> 1974</u>	<u> 1976</u>	<u> 1978</u>	<u>1980</u>	<u> 1982</u>	<u>1984</u>	<u> 1986</u>	<u> 1988</u>	<u>1990</u>	<u> 1992</u>	<u> 1994</u>	<u> 1996</u>	<u> 1998</u>	<u>2000</u>
Black	061*	109***	063*	033	105***	.102***	.161***	.174***	.101***	.058*	026	.024	.211***	.263***	321***
Yrs. of Education	.228***	.190***	.245***	.206***	.249***	.211***	.232***	.195***	.283***	.211***	.256***	.238***	.254***	.212***	.277***
Male	035*	029	093***	069**	126***	060***	171***	073	143***	089***	131***	041**	153***	060***	163***
Married	.392***	.517***	.443***	.532***	.528***	.538***	.460***	.541***	.604***	.498***	.418***	.538***	.486***	.535***	.550***
Employed	.291***	.331***	.302***	.266***	.319***	.283***	.261***	.269***	.400***	.260***	.311***	.266***	.229***	.221***	.262***
Age	.031***	.039***	.038***	.041***	.039***	.042***	.034***	.041***	.045***	.039***	.031***	.042***	.036***	.041***	.035***
Constant	-3.753***	-4.653***	-4.384***	-4.763***	-4.464***	-4.754***	-4.149***	-4.776***	-5.615***	-4.904***	-4.265***	-5.368***	-4.750***	-5.181***	-4.945***

	<u> 1972 </u>	1976	1980	1984	1988	1992	1996	2000
Black	2.553***	2.824***	3.066***	2.375***	2.307***	2.471***	3.241***	2.373***
Yrs. of Education	006	067***	064*	015	004	004	076**	.033
Male	244	039	358*	151	162	270*	439***	430**
Married	163	.049	136	070	021	439***	223	217
Employed	026	.197	.335	001	.154	.078	.411*	.137
Age	019***	010*	.001*	.000	003	.007	.004	.000
Income	128*	276***	123	303***	191*	173**	160*	153*
Constant	.723*	1.817***	.058*	716	.916*	.354	1.423**	307

SENATE VOTE ANALYSIS (voted for Democratic Candidate = 1)

	<u> 1972</u>	1974	1976	1978	1980	1982	1984	1986	1988	1990	1992	1994	1996	1998	2000
Black	1.719***	2.289**	1.852***	1.392**	1.525***	1.729***	1.930***	1.915***	2.049***	2.052***	1.946***	1.029***	2.072***	1.719***	1.906***
Yrs. of Education	041	026	027	081**	027	088*	041	062	068*	039	048	023	030	042	021
Male	010	184	151	.706***	.145	028	076	050	130	156	239	513**	349	061	237
Married	.017	.393	.117	215	157	.079	.262	163	043	106	339*	081	569**	.155	128
Employed	.218	.516*	090	.236	116	.179	.076	117	.195	.082	057	.110	.076	.291	.047
Age	015**	014*	011*	007	.005	003	001	003	005	002	.001	.008	.003	.008	.001
Income	260***	319**	116	253*	077	123*	237*	.022	195*	134	035	187*	230*	038	115
Constant	1.673***	1.780**	1.589**	1.763**	.482	1.672*	1.092	1.010	1.894***	1.352	1.054*	.409	1.475*	.117	.876

^{*}p < .05 **p < .01 (two-tailed tests)